Gibrat’s law in Brazilian franchising: an empirical note

Marcelo Resende  
*Instituto de Economia, Universidade Federal do Rio de Janeiro*

Vicente Cardoso  
*Banco Nacional de Desenvolvimento Econômico e Social-BNDES*

This paper can be downloaded without charge from http://www.ie.ufrj.br/
Gibrat’s law in Brazilian franchising: an empirical note

August, 2012

Marcelo Resende
Instituto de Economia, Universidade Federal do Rio de Janeiro,
Av. Pasteur 250, Urca, 22290-240, Rio de Janeiro-RJ
mresende@ie.ufrj.br

Vicente Cardoso
Banco Nacional de Desenvolvimento Econômico e Social-BNDES
Av. República do Chile 100, 20031-917, Rio de Janeiro-RJ, Brazil
vicentescardoso@bndes.gov.br

Abstract
The paper investigates the lognormality of firm size distribution in the context of the Brazilian franchising. This implication of Gibrat’s law-GL is considered in a yearly basis under two settings. The evidence for both the totality of firms and for mature firms at least 5 years old was generally consistent with GL. However, in the latter case moderate departures from lognormality are observed in some years.

Keywords: Gibrat’s law; franchising; nonparametric kernels

JEL Classification: L11; L26.
1 Introduction

Gibrat’s law-GL is a recurring topic in the literature pertaining firm growth and dynamics and the renewed interest is documented in Sutton (1997). The assumption of independence between firm growth and size has been frequently challenged in the empirical literature on the basis of the underlying conceptual aspects and yet the increasing data availability and more careful statistical and econometric assessments [see Santarelli et al. (2006) for a comprehensive survey of the empirical literature].

As a rule, the literature focused on the manufacturing industry in developed countries and often detects a negative relationship between firm growth and size and age, where Evans (1987) and Hall (1987) are representative studies for the U.S.. It is worth mentioning that scale aspects might play an important role in the rejection of the law for the manufacturing industry. In that sense, Audretsch et al. (2004) advanced the possibility that services industries could provide a more favorable setting for the validity of GL and indeed the evidence for Dutch firms in the hospitality industries was encouraging. However, Piergiovanni et al. (2003) studied Italian new-born firms in that segment and obtained support for the GL only in 2 out 5 of the business groups considered.

As stated by Sutton (1997), GL comprises two assumptions, being the first that the “next opportunity is taken up by any particular active firm is proportional to the current size of the firm” and the second that firm growth should be independent of size. The second condition will follow from the first, as mentioned by Audretsch (2002), only if size is not related to survival. This is because if growth is random but proportional to firm size, then the growth rates should be equal on average. But if size influences the chances of survival, it follows that GL will not hold in a sample with small and large firms, whereas it will if only larger firms are considered. Following this reasoning, in industries where economies of scale are absent and sunk costs are not relevant, there is not a theoretical case to expect that smaller firms would have a lower survival probability (due to higher costs) than their larger counterparts, hence growth rates tend to be closer to independent of size.

In general, the assessment of GL in service industries provide at most some partial support in a few cases of the handful of studies conducted so far. Lotti (2007) detects
significant associations between firm growth and size in selected sectors of the Italian service industries. Hardwick and Adams (2002), on the other hand, focused on the insurance industry in the U.K. and found supportive evidence for GL in the long run, though violations were observed for shorter time intervals.

Finally, Maçãs Nunes and Serrasqueiro (2009) considered the service sector in Portugal and found a negative relationship between firm growth and size which indicated an important role for ownership control in that context.

The present paper intends to investigate distributional regularities implicated by GL in the context of the Brazilian franchising segment and different motivations can be evoked:

a) Previous assessments of GL concentrated on developed countries;

b) The small literature on GL in the case of service industries could further benefit from the study of the franchising segment. Quantitative studies for that sector in developing economies are not common [Façanha et al. (2013) provide an exception in terms of the investigation of firm survival in Brazilian franchising] and it provides a potentially favorable setting for the prevalence of GL as small scaled business can prosper and scale gains are likely to prevail mostly in the centralized provision of inputs and training.

The paper is organized as follows. The second section discusses conceptual issues associated with Gibrat’s law. The third section discusses the data base. The fourth section presents the empirical results. Finally, the fifth section brings some final comments.

2 Gibrat’s law and distributional regularities

Beyond direct assessments of the firm growth and size relationship, the empirical literature has discussed distributional regularities that could emerge and possible generating mechanisms. Synthetic road maps are presented in Vining (1976) and Resende (2004b). GL contends that the probability of a given proportionate change in size (during a particular period) is the same for all firms in a given industry independent
of their size at the beginning of the time period [see e.g. Mansfield (1987)]. The usual argument is presented, for example, in Kalecki (1945), Saboia (1977) and Hay and Morris (1991).

Let \( S_t \) denote the size of a given firm in period \( t \) and let \( \varepsilon_t \) stand for the growth rate of the form relative to the previous period, then it follows that:

\[
S_t = S_0 (1 + \varepsilon_1) \quad \text{and after recursive substitutions one obtains:}
\]

\[
S_t = S_0 (1 + \varepsilon_1)(1 + \varepsilon_2)...(1 + \varepsilon_t) \quad \tag{1}
\]

Moreover, let \( Y_i = \log S_i \) for \( i=0, t \) and \( y_i = \log(1+\varepsilon_i) \) for \( i=1,2...t \). Taking the logarithm of expression (1), it follows:

\[
Y_t = Y_0 + y_1 + y_2 +...+ y_t \quad \tag{2}
\]

Considering a first-order Taylor expansion around zero, one obtains:

\[
Y_t \approx Y_0 + \varepsilon_1 + \varepsilon_2 +...+ \varepsilon_t \quad \tag{3}
\]

where one is using the approximation \( \log(1+\varepsilon_i) \approx \varepsilon_i \) for \( i=1,2,...,t \). Assuming that the growth rates are independent of the initial firm size and that this has finite mean \( \mu_\varepsilon \) and variance \( \sigma_\varepsilon^2 \), it is possible to consider a Central Limit Theorem and conclude that the distribution of \( Y_t \) can be approximated by a normal distribution with mean 0 and variance 1 as \( t \to \infty \). Therefore one can consider the log-normal distribution for firm size as a long run implication of GL.\(^1\) That distributional regularity is robust even when one allows for negative correlation between firm growth and size [see Kalecki (1945)] or consider a more general autocorrelation structure in terms of an ARIMA model [see Saboia (1977)].

The next generation of stochastic growth models included Simon (1955), Simon and Bonini (1958), Ijiri and Simon (1964) and Steindl (1965), and pinpointed the emergence of Pareto and Yule distributions when one allows entry and exit dynamics. Therefore,

\(^1\) Asymptotically the contribution of the initial firm size \( Y_0 \) would be negligible as \( t \to \infty \). In the general case a somewhat more complex expression would arise.
the independence assumption underlying GL leads to skewed distributions under distinct hypotheses. Cabral and Mata (2003) investigate shapes of the firm size distribution in Portugal for unconditional and conditional cases what can be relevant for considering some aspects emphasized by the literature as for example the age of the firm.\footnote{Resende (2004b) implemented tests for lognormality of firm size at the sectorial level in the Brazilian manufacturing industry and a strong rejection of that implication of GL was indicated.} Simon and Bonini (1958) outline some possible factors that could lead to violations in GL as for example: i) non growth objectives, ii) merger activity, iii) new investments, iv) regional demand, v) aggressiveness in marketing and management, vi) customer brand preference, vii) firm age.

As for the growth and size relationship it is important to identify different constraints to firm expansion that could be directly related and favored by a larger scale. In fact, the financial constraints for growth in the case of the franchising segment possess a distinct character as the investment in stores is enabled by means of the fixed component of the contract (the franchise fee).

Michael (1996) contends that economies of scale are likely to arise in connection with marketing, purchasing and product development. In fact, nation-wide TV advertisements are usual for some mature fast food chains. Nevertheless, the existence of potential agency problems in terms of moral hazard issues can put quality standards at risk and requires costly monitoring activities. Those aspects can counteract scale gains and tend to be more complex if the degree of heterogeneity across franchisees is substantial. Moreover, independent of the level of effort in providing quality, different locations offer distinct degrees of risk expressed, for example, in terms of the variability of sales [see related discussion in Martin (1988)].

The franchising business format could in principle lighten the importance of scale advantages in general and make the independence between firm growth and size more tenable when those potential scale gains are not particularly salient. However, for newly created firms it still would be possible to observe expansion constraints reflecting a brand that is not yet consolidated. In any case, it would be relevant to also consider age aspects in empirical analyses of the franchising segment.
3 Data sources

The Brazilian franchising association (Associação Brasileira de Franchising-ABF) conducts an annual survey published in the so-called Guia das Franquias, where detailed data from the previous year is collected with respect to different aspects of the contract (franchise fee, royalty fee, advertising fee among others), sector of activity, date of foundation and different qualitative information. This paper considers the 1994-1999 period (available in the annual reports from 1995 until 2000) so as to assure the homogeneity of the data. Indeed, up to the beginning of the 1990s, that data source included also contracts that could not be characterized as a typical franchising scheme as for example brand licensing agreements and after 2000 (annual report in 2001) the publication became less comprehensive as it started to exclude firms that were not associated to ABF. The firm size variable is obtained by multiplying the average store area (in squared meters) by the number of outlets. It is reasonable to assume some proportionality with the total number of employees that was not reported. In fact, the layout of stores within a chain tend to be somewhat standardized under franchising. The analysis is developed upon the natural log of that variable.

The minimum number of firms was 478 in 1997 whereas the maximum number of firms was 720 in 1995. Further analysis was carried out for more mature firms with a smaller sample. In the case of firms with at least 5 years since foundation the sample sizes ranged from 128 to 256 firms in different years. Those reduced sample sizes reflect inconsistent age reporting that required smaller samples in order to conduct a reliable analysis and yet retain acceptable asymptotic properties for the nonparametric estimation. Unfortunately, additional analysis in terms of a longer horizon with firms with at least 10 years since foundation was not feasible given the small samples thus obtained.
4 Empirical analysis

The lognormality issue is approached by means of kernel estimators for the density function. All the analyses were carried out with Stata 12.0. We consider 2 levels of analysis:

a) Lognormality analysis for the totality of firms in each year;

b) Lognormality analysis for the totality of firms in each year filtered by a minimum age in general (based on the foundation date).

It is worth mentioning that we use the Epanechnikov kernel function. In fact, the really critical choice concerns the window width for which we adopt Silverman’s rule of thumb.3

The initial case is considered in figure 1.

3 That rule adopts a criterion based on the minimization of the integrated mean squared error, see Pagan and Ullah (1999) for details.
Figure 1

Kernel density estimation – Brazilian Franchising segment 1994-99 – totality of firms
The evidence appears to favor the lognormal shape implied by GL for all the years studied. However, when one considers the distributions in terms of mature firms with at least 5 years since foundation and from the start of franchising activities. The corresponding results appear respectively in figure 1 and 2.

**Figure 2**

*Kernel density estimation – Brazilian Franchising segment 1994-99 – firms with at least 5 years since the foundation date*
One can note some moderate discrepancies with lognormality when one controls for age but still that implication of GL appears to prevail in franchising. We had access to data on years of franchising experience. However, the smaller size samples in that case would not provide sufficient statistical rigor in that case despite the similar evidence obtained for that age criterion.

5 Final comments

The paper aimed at testing the lognormal distributional shape that is implied by Gibrat’s law-GL. For that purpose we considered the totality of firms along the 1994-99 period on a yearly basis and sub-samples with mature firms with at least 5 years of existence.

The support of GL was very strong for the totality of firms and for more mature ones, though in the former case some moderate departures are observed for some periods.

Altogether, in some years the departures are moderate so that it appears that indeed franchising provides a more favorable setting for testing that regularity vis-à-vis industrial firms. Scale gains (especially in advertising) are likely to be associated with larger mature firms while large scale marketing initiatives tend to prevail only in selected sectors as for example fast food and cosmetics. Nevertheless, the existing departures are likely to reflect some partial relevance for scale aspects that are not readily observable and relate to centralized training and purchasing of inputs, for example. Additional investigations that consider the role of franchising experience should be carried out in the future should the necessary data become available.
References


